Electronic Supplementary Information for 'Comment on "Aqueous SET-LRP catalyzed with in situ generated Cu(0) demonstrates surface mediated activation and bimolecular termination" by S. Samanta et al., Polym. Chem. 2015, 6, 2084'

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Details of non-linear least squares fitting procedure.

Conversion and residual functionality (f_{PBF}) data were read from Fig. 8b of reference 1, and are shown in Table S1.

Table S1. Conversion and residual functionality (f_{PBF}) data from Fig. 8b of reference 1, and results of NLLS fitting assuming error in both variables with k_t/k_p^2 = 0.047.

^a A value of 0.999 was used to represent 100% conversion.

The data were fitted to the following single-parameter model:

$$
\hat{f}_{PBr} = \frac{4.61}{60 \times 0.09} \times a \ln(1 - conv)
$$

Equivalently,

$$
c\widehat{on}v=1-e^{\frac{60\times0.09}{4.61}\frac{f_{PBr}}{a}}
$$

In each equation, a is the parameter to be fitted and represents the value of k_p/k_t^2 . Using these equations, fitted values of conversion and f_{PBF} can be obtained from experimental values of f_{PBR} and conversion, respectively.

The fitting procedure is carried out with the aim of minimizing the sum of squared residuals, $\sum r^2$, assuming an error of similar magnitude in both variables. The distance, r, between the model curve and a data point (x,y) (Figure S1) is approximately given by:

$$
r^2 \approx \frac{(\Delta y)^2}{1 + \left(\frac{\Delta y}{\Delta x}\right)^2}
$$

This is equivalent to carrying out a conventional non-linear least squared fit, assuming negligible error in

conversion, with the residuals weighted by a factor of $\frac{1}{\sqrt{2}}$ $1+\left(\frac{\Delta y}{4\mu}\right)$ $\frac{\Delta y}{\Delta x}$.

Figure S1. Relationship between distance between experimental data point and model curve, *r*, and distances between experimental and fitted data points, Δx and Δy .

Fitting was carried out by an iterative procedure, in which an initial estimate of *a* was used to generate a new estimate by solving the matrix equation

$$
\Delta a = \left(\mathbf{J}^{\mathrm{T}} \mathbf{W} \mathbf{J}\right)^{-1} \mathbf{.} \mathbf{J}^{\mathrm{T}} \mathbf{W} \Delta \mathbf{y}
$$

Where **J** is the Jacobian matrix whose elements are given by

$$
J_i = \left(\frac{\partial \hat{f}_{PBr}}{\partial a}\right)_i
$$

W is the weight matrix

$$
W_{ii} = \frac{1}{1 + \left(\frac{\Delta y}{\Delta x}\right)_i^2}
$$

$$
W_{ij} = 0 \ (i \neq j)
$$

and **Δy** is the vector of residuals

$$
\Delta y_i = f_{PBr,i} - \hat{f}_{PBr,i}
$$

The new value of *a* was obtained by adding Δ*a* to the previous value until the difference between iterations was less than 1 \times 10⁻⁵. This procedure gave a best estimate for *a* of 0.0468 s·mol·L⁻¹.

The standard error in *a* can be estimated from

$$
se_a = \sqrt{\frac{(\mathbf{J}^{\mathsf{T}} \mathbf{W} \mathbf{J})^{-1} \Delta \mathbf{y}^{\mathsf{T}} \mathbf{W} \Delta \mathbf{y}}{n-1}}
$$

where n is the number of observations (5 in this case).

The standard error in *a* was thus estimated at 0.0133 s·mol·L⁻¹. Figure S2 shows the curve of best fit $(k_p/k_t^2 = 0.0468)$ together with the curves corresponding to $k_p/k_t^2 = 0.0468 \pm 0.0133$. The probability that the true value of k_t/k ² falls in this range is approximately 70%.

Figure S2. Curve of best fit (solid line) to experimental residual functionality and conversion data from reference 1, Fig. 8b, with curves corresponding to best estimate of k_t/k_p^2 plus or minus the standard error (dashed lines).

Figure S3. Values of k_p and k_t which are consistent with the experimental data shown in Figure S2. Solid line indicates values of k_p and k_t which satisfy k_t/k_p^2 = 0.0468. Dashed lines correspond to k_t/k_p^2 = 0.0468 \pm 0.0133 (best estimate \pm 1 standard error). The probability that the true values of k_p and k_t fall between the dashed lines is approximately 70%.