

Supplementary Information

Efficient and robust image registration for two-dimensional micro-X-ray fluorescence measurements

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A Error estimation for graph registration

We show here that, for GRAPH REGISTRATION with a Moore graph, $\text{Var}[\hat{s}_j]$ tends to 0 as the parameter d of the Moore graph tends to infinity. We assume here a bit more familiarity with graph theory concepts than in the rest of the article. All concepts can be found in a standard textbook.¹

A Moore graph is a d -regular graph with diameter k , where the number of vertices is $n = 1 + d \sum_{i=0}^{k-1} (d-1)^i$. The structure of the Moore graph ensures that the union of paths $P_{i,j}$ that end in j form a balanced tree rooted in j . That is, there are $d(d-1)^{\ell-1}$ vertices at depth ℓ and each of them have $\frac{n-1-d \sum_{m=0}^{\ell-1} (d-1)^m}{d(d-1)^{\ell-1}}$ descendants. As a consequence, if we consider the shortest paths between a fixed vertex j and all other vertices then an edge between a vertex at depth ℓ and its parent appears as often as:

$$1 + \frac{n-1-d \sum_{m=0}^{\ell-1} (d-1)^m}{d(d-1)^{\ell-1}} = \frac{n-1-d \sum_{m=0}^{\ell-2} (d-1)^m}{d(d-1)^{\ell-1}}$$

¹Diestel, R. Graph Theory (5th edition). Springer-Verlag, 2017.

With the same assumptions on the errors as before, we compute:

$$\begin{aligned}
\text{Var} [\hat{s}_j] &= \text{Var} \left[\frac{1}{n} \sum_{i \neq j} \sum_{(u,v) \in E(P_{i,j})} \varepsilon_{u,v} \right] \\
&= \text{Var} \left[\frac{1}{n} \sum_{\ell=1}^k \sum_{\substack{\text{vertex } i \\ \text{at depth } \ell}} \left(\frac{n-1-d \sum_{m=0}^{\ell-2} (d-1)^m}{d(d-1)^{\ell-1}} \right) \varepsilon_{i,\text{parent}(i)} \right] \\
&= \frac{1}{n^2} \sum_{\ell=1}^k \sum_{\substack{\text{vertex } i \\ \text{at depth } \ell}} \left(\frac{n-1-d \sum_{m=0}^{\ell-2} (d-1)^m}{d(d-1)^{\ell-1}} \right)^2 \text{Var} [\varepsilon_{i,\text{parent}(i)}] \\
&\leq \frac{1}{n^2} \sum_{\ell=1}^k d(d-1)^{\ell-1} \left(\frac{n-1-d \sum_{m=0}^{\ell-2} (d-1)^m}{d(d-1)^{\ell-1}} \right)^2 \varepsilon \\
&= \frac{1}{n^2} \sum_{\ell=1}^k \frac{\left(n-1-d \sum_{m=0}^{\ell-2} (d-1)^m \right)^2}{d(d-1)^{\ell-1}} \varepsilon.
\end{aligned}$$

We continue

$$\text{Var} [\hat{s}_j] \leq \frac{1}{n^2} \sum_{\ell=1}^{\infty} \frac{n^2}{(d-1)^\ell} \varepsilon = \frac{1}{d-2} \varepsilon,$$

where in the last step we have used the identity of a geometric series. Therefore, the variance decreases with d , and thus, with high probability and increasing d , the estimate \hat{s}_j will be close to the true value s_j .

B Expected length of a permutation path

We compute the expected length ℓ of a path of a permutation τ of the positions. The computation involves three expectations: the expectation $\mathbb{E}_{s_1, \dots, s_n}$ with respect to the random process to draw the shifts s_i , the expectation $\mathbb{E}_{s_{1,2}, \dots, s_{k,k+1}}$ to draw the shifts between consecutive measurements (governed by a 2-dimensional normal distribution with mean 0 and variance σ^2 in each direction), and the expectation \mathbb{E}_τ with respect to the choice of the random permutation τ . We furthermore use \mathbb{P} to denote a probability.

$$\begin{aligned}
\mathbb{E}_{s_1, \dots, s_n} \mathbb{E}_\tau [\ell] &= \mathbb{E}_{s_1, \dots, s_n} \mathbb{E}_\tau \left[\sum_{i=1}^{n-1} \|s_{\tau(i+1)} - s_{\tau(i)}\| \right] \\
&= \sum_{i=1}^{n-1} \mathbb{E}_\tau \mathbb{E}_{s_1, \dots, s_n} [\|s_{\tau(i+1)} - s_{\tau(i)}\|] \\
&= \sum_{i=1}^{n-1} \sum_{k=1}^{n-1} \mathbb{P} [|\tau(i+1) - \tau(i)| = k] \mathbb{E}_{s_{1,2}, \dots, s_{k,k+1}} \left[\left\| \sum_{j=1}^k s_{j,j+1} \right\| \right] \\
&= \sum_{i=1}^{n-1} \left[\sum_{k=1}^{n-1} \mathbb{P} [|\tau(i+1) - \tau(i)| = k] \sqrt{\frac{\pi}{2}} \sqrt{k\sigma^2} \right],
\end{aligned}$$

where the last step uses that $\|\sum_{j=1}^k s_{j,j+1}\|$ is distributed according to a Rayleigh distribution² with scale parameter $\sqrt{k\sigma^2}$. We continue

$$\begin{aligned}\mathbb{E}_{s_1, \dots, s_n} \mathbb{E}_\tau [\ell] &= (n-1) \sqrt{\frac{\pi}{2}} \sigma \sum_{k=1}^{n-1} \frac{(n-k)}{n(n-1)} \sqrt{k} \\ &= \frac{\sqrt{2\pi}\sigma}{2n} \sum_{k=1}^{n-1} (n-k) \sqrt{k} \\ &= \frac{\sqrt{2\pi}\sigma}{2} \left(\sum_{k=1}^{n-1} \sqrt{k} - \frac{1}{n} \sum_{k=1}^{n-1} k^{\frac{3}{2}} \right)\end{aligned}$$

Using the Euler-Maclaurin formula³ for both sums, this can be written as:

$$\begin{aligned}\mathbb{E}_{s_1, \dots, s_n} \mathbb{E}_\tau [\ell] &= \frac{\sqrt{2\pi}\sigma}{2} \left(\left(\frac{2}{3} n^{\frac{3}{2}} + \frac{1}{2} n^{\frac{1}{2}} + O(1) \right) - \left(\frac{2}{5} n^{\frac{3}{2}} + \frac{1}{2} n^{\frac{1}{2}} + O(1) \right) \right) \\ &= \frac{2\sqrt{2\pi}}{15} \sigma n^{\frac{3}{2}} + \sigma O(1)\end{aligned}$$

(The $O(1)$ -notation means: a quantity that is bounded as n grows to infinity.)

²Papoulis, A. and Pillai, S. U. Probability, random variables, and stochastic processes. Tata McGraw-Hill Education, 2002.

³Apostol, T. M. An elementary view of Euler's summation formula. *The American Mathematical Monthly*, 106(5): 409–418, 1999.